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Asymptotic dependence and asymptotic independence

General: $P(X \leq x, Y \leq y) = \exp\left[-\left(\frac{1}{x} + \frac{1}{y}\right) A\left(\frac{x}{x+y}\right)\right]$

Richard's fit.

$$P(X > x, Y > x) \sim \frac{2(1 - A(\frac{1}{2}))}{x} \text{ unless } A(\frac{1}{2}) = 1$$

$$\chi = \lim_{x \rightarrow \infty} P(Y > x | X > x) = 2 \left\{ 1 - A\left(\frac{1}{2}\right) \right\}$$

Aside: assumes unit Fréchet margins for both X and Y . If not, need to transform, but same conclusion.

$\chi = 0$: asymptotically independent

$\chi > 0$: asymptotically dependent

In this case, $\chi = 0$ means X and Y are fully independent, but this may not be true in general.

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Example (Ledford & Tawn 1996)

If $(X, Y) \sim \text{Normal}$, correlation ρ , then
 $P(X > x, Y > x) \sim C x^{-\frac{2}{1+\rho}} (\log x)^{-\frac{\rho}{1+\rho}}$

$\chi = 0$ but dependence is nontrivial

New concept!

$$P(X > x, Y > x) \sim L(x) x^{-1/\eta} \quad 0 < \eta \leq 1$$

η : "coefficient of tail dependence"

L is slowly varying: $\frac{L(tx)}{L(x)} \rightarrow 1$ as $x \rightarrow \infty$
for all $t > 0$.

Normal case: $\eta = \frac{1+\rho}{2}$

$\frac{1}{2} < \eta < 1$: positive association

$0 < \eta < \frac{1}{2}$: negative

$\eta = \frac{1}{2}$: "near independent"

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To estimate η : define $T = \min(X, Y)$

$$P(T > u+t | T > u) \sim \left(1 + \frac{t}{u}\right)^{-1/\eta} \quad (u \rightarrow \infty)$$

Fit GPD to estimate η (Ledford-Tawn 1996)

Extension to full joint tail (L&T 1997)

Proposed
$$P(X > x, Y > y) = \mathcal{L}(x, y) x^{-c_1} y^{-c_2}$$

where $c_1 + c_2 = \frac{1}{\eta}$, \mathcal{L} is bivariate SV in sense

$$g(x, y) = \lim_{t \rightarrow \infty} \frac{\mathcal{L}(tx, ty)}{\mathcal{L}(t, t)} \text{ exists}$$

$$g(cx, cy) = g(x, y) \quad \forall c > 0, \text{ so } g(x, y) = g_*\left(\frac{y}{x+y}\right)$$

some ~~g~~ g_*

L&T proposed
$$P(X > x, Y > y) \sim K g_*\left(\frac{y}{x+y}\right) x^{-c_1} y^{-c_2}$$

but no clearly defined model for g_*

Comment on terminology: both η and \mathcal{L} have been called "coefficient of tail dependence". Confusing!

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Romance Ledford (2009, 2011)

Assume $P(X > x, Y > y) = \frac{I(x, y)}{(xy)^{1/\alpha}}$

$$\frac{I(ux, uy)}{I(x, y)} \rightarrow g(x, y) \text{ as } u \rightarrow \infty$$

$$[\equiv C_1 = C_2 \text{ in L&T theory}]$$

R&L define (S, T) as weak limit of $(\frac{X}{u}, \frac{Y}{u})$

given $X > u, Y > u$. Then

$$\bar{F}_{ST}(s, t) = P(S > s, T > t) = \frac{g_+(s/(s+t))}{(st)^{1/\alpha}}$$

$$\text{Let } R = S + T, W = \frac{S}{R}$$

$$\mu_{RW}(dr, dw) = r^{-(1+1/\alpha)} d\mathbb{H}_\alpha(w)$$

\mathbb{H}_α some non-neg measure on $[0, 1]$

leads to eqn. (4.55) p. 130

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$$\text{Example: } F_{xy}(x, y) = \frac{\lambda u^{1/\eta}}{N_p} \left[(px) \overset{-1/\eta}{+} \left(\frac{y}{p} \right) \overset{-1/\eta}{+} \right. \\ \left. - \left\{ (px)^{-1/\alpha} + \left(\frac{y}{p} \right)^{-1/\alpha} \right\}^{\alpha/\eta} \right]$$

Parameters α, ρ, η , asy. dep if $\eta \geq \alpha$, asy. ind. if $\eta < \alpha$

Extension to $p > 2$: Qin, Smith & Ren (2009),
Renus-Ledford (2011).

Applications: example by RLS discussing papers
by Herweiger & Seager (2008), Lund Kim (2012)

Figures on p. 132, 133; Tables on p. 134

Other approaches to MV extremes:

(a) Heffernan-Tawn approach (2004): Focus on
limit distribution of $(Y_1, \dots, Y_i, Y_{i+k}, \dots, Y_p)$

given $Y_i \rightarrow \infty$.

(b) High-dim MV dist's: Dan Cooley